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Fast and effective: Intuitive processes in complex decisions

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Abstract

Is it possible to carry out complex multi-attribute decisions (which require an estimation of the weighted average) intuitively, without resorting to simplifying heuristics? Over the course of 600 trials, 26 participants had to choose the better-suited job-candidate, a task requiring comparison of two alternatives over three/four/five dimensions with specified importance weights, with a time constraint forcing intuitive decisions. Participants performed the task fast (mean reaction time (RT) ~ 1.5 s) and with high accuracy (~86%). The participants were classified as users of one of three strategies: Weighted Additive Utility (WADD), Equal Weight rule and Take-The-Best heuristic (TTB). Fifty-nine percent of the participants were classified as users of the compensatory WADD strategy and 29% as users of the non-compensatory TTB. Moreover, the WADD users achieved higher task accuracy without showing time costs. The results provide support for the existence of an automatic compensatory mechanism in weighted average estimations.

Keywords Decision making · Weighted average · Compensatory process · Take-the-best

Introduction

Complex decisions, such as selecting a job candidate or a vacation package, are among the most demanding and challenging human activities. A major cause of this challenge is the presence of trade-offs between attributes (e.g., intelligence vs. motivation for job candidates) that are difficult to compare. While a normative theory, based on weighted additive utility (WADD), was developed by early decision theorists (Keeney & Raiffa, 1976), a widely accepted view considers that the computations required for the normative WADD algorithm are too complex for online human decisions (not assisted by offline calculations or external aids). Accordingly, it is often

assumed that when faced with such decisions, humans typically resort to a number of simplifying non-compensatory heuristics, such as Take-the-Best (TTB), according to which one chooses on the basis of the most important attribute (in case of a tie, the second most important attribute is considered; Gigerenzer & Goldstein, 1996, 1999; Payne, Bettman, & Johnson, 1993; but see Newell, 2005, for a critique of this approach and suggestions of formal models of ecological rationality). Such heuristics simplify the decision algorithm, by replacing the compensatory processes – in which all the attributes are weighted into the decision – with a non-compensatory one, in which only a small subset of the attributes is taken into account (Dieckmann & Rieskamp, 2007; Gigerenzer & Goldstein, 1999; Tversky, 1969, 1972).

Recent research has challenged the assumption that compensatory strategies are too complex and thus beyond daily decision-making ability. First, numerous studies in the domain of probabilistic inference with binary cues have shown that even when environments are designed to promote the use of TTB heuristic, a significant proportion of participants do not “take the best” (e.g., Bröder, 2000; Lee & Cummins, 2004; Newell & Shanks, 2003). Second, more recent experimental work has demonstrated that most participants make probabilistic inferences based on multiple cues in a compensatory yet rapid and automatic manner (Glöckner & Betsch, 2008, 2012; Glöckner, Hilbig, & Jekel, 2014). Other research has

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60 manipulated time pressure, confirming the presence of compen-
 61 satory strategies with a 3-s response-deadline and, for
 62 some participants, even for a strict deadline of 750 ms (Oh
 63 et al., 2016).

64 A mechanistic account of such an automatic yet compen-
 65 satory decision process was proposed by Glöckner and col-
 66 leagues in the form of the *PCS* model (Glöckner et al., 2014).
 67 *PCS* is a connectionist, accumulator-type model, that inte-
 68 grates (using a parallel architecture) differences in weighted
 69 evidence (ΔWA) between the alternatives and predicts slower
 70 reaction times (RTs) for decisions with smaller ΔWA
 71 (Glöckner & Betsch, 2008, 2012; see also Roe, Busemeyer,
 72 & Townsend, 2001 for a *Decision-Field-Theory* model of
 73 multi-attribute decisions).

74 In this paper we demonstrate an ability to make complex
 75 decisions using compensatory, rapid, and automatic mecha-
 76 nisms in a different domain: multi-attribute decision-making
 77 based on numerical (non-binary) attributes. Such decisions
 78 normatively require a weighted averaging computation, tradi-
 79 tionally associated with analytical processes. Moreover, multi-
 80 attribute decisions with non-binary attributes have received
 81 less attention in recent research (see Russo & Doshier, 1983
 82 and Tversky, 1969 for some older studies) and they differ from
 83 binary cue decisions in a number of important aspects (prob-
 84 lem-space is virtually infinite and precludes the use of simpli-
 85 fying strategies, such as memorizing or counting). Therefore,
 86 if rapid and compensatory (*WADD*) strategies can be deployed
 87 in this domain, this would provide support for the impressive
 88 power of the intuitive decision-maker.

89 Recent research has shown that an important precursor of
 90 *WADD* – numerical averaging – can be estimated in a relative-
 91 ly precise and yet automatic manner (Brezis, Bronfman,
 92 Jacoby, Lavidor, & Usher, 2016; Brezis, Bronfman, &
 93 Usher, 2015; Rusou, Zakay, & Usher, 2017). Here we set to
 94 test whether this ability extends to weighted averaging, by
 95 employing a job selection multi-attribute decision task.

96 **Experiment**

97 Participants were asked to take the role of a job interviewer
 98 who chooses one of two candidates based on the candidates’
 99 abilities on several attributes and their relative importance. We
 100 varied (in blocks) the number of attributes (three/four/five),
 101 and we presented a large set of decision problems with ran-
 102 domized values (see Methods). This design allows us to con-
 103 trast decision strategies within each participant using choices
 104 and RTs. While the *TTB* heuristic predicts slower decisions in
 105 cases there is a tie on the most important attribute, *PCS* (or
 106 other accumulator models) predicts decision times that in-
 107 crease with lower ΔWA . Another central question of interest
 108 is whether the deployment of compensatory strategies results
 109 in improved task performance.

Method

110

Participants

111

112 Twenty-six students from Tel-Aviv University (14 females,
 113 age: 19–31, $M=24.7$) participated in the experiment, in ex-
 114 change for course credit and payment that was dependent on
 115 performance. On average, participants received 30 NIS (~7.5
 116 USD). The sample size was set at 26 with each subject tested
 117 in three tasks, allowing for 78 (26×3) classifications in total
 118 to be made (see *Strategy Classifications* section).

Materials

119

120 Each decision was presented in a table-format (see Table 1).
 121 Three jobs were presented, with three, four, and five attributes,
 122 respectively. Each job specified the attributes' importance (i.e.,
 123 weight; see Table 1). When the job had three attributes, the
 124 specified importance-weights were 3, 2, and 1 (i.e., the most
 125 important attribute was three times more important than the
 126 least important attribute), for the four-attribute job they were
 127 4, 3, 2, and 1, and for the five-attribute job 5, 4, 3, 2, and 1.
 128 The values the candidates received in each trial were generat-
 129 ed randomly, as random integers between 1 (poor) and 9 (ex-
 130 cellent; from a uniform distribution; if the resulting weighted
 131 average for the two candidates was the same, the ratings were
 132 generated anew).

133 A time limit for providing an answer was imposed, in order
 134 to encourage participants to rely on their intuitive mind-set
 135 and not explicitly compute weighted averages (Horstmann,
 136 Hausmann, & Ryf, 2010). The time limit increased with the
 137 number of attributes that had to be considered to make sure
 138 that all the information can be encoded. The time limits were
 139 3 s for the 3-attributes, 4 s for the 4-attributes, and 5 s for the 5-
 140 attribute jobs. As we report below, however, the time con-
 141 straints did not affect the actual decision times.

Procedure

142

143 Participants completed 600 trials overall, with three
 144 blocks of 200 trials for each job (3, 4, or 5 attributes).
 145 On each trial, a choice problem (see Table 1) was pre-
 146 sented until the participant entered a decision by using
 147 the keyboard. Visual feedback (correct/incorrect) was
 148 given after each trial, based on the weighted averages.
 149 Feedback was also given on the number of correct trials
 150 the participant accumulated, which was translated to
 151 monetary reward at the end of the experiment. Once the
 152 time limit expired, the trial ended, and the visual feed-
 153 back, (“too slow,” was presented. The whole procedure
 154 took approximately 60 min (see [Supplementary Material](#)
 155 for details).

t1.1 **Table 1** Example of a trial in which the job had four attributes, with weights of 4, 3, 2, and 1. Here candidate A had the higher weighted average (5.2 vs. 4.3 for candidate B) and so she should be selected for the job, while candidate B should be selected according to *TTB* heuristic

t1.2	A	B
t1.3 Intelligence – 4	3	3
t1.4 Work Ethic – 3	5	7
t1.5 Easy to Work With – 2	9	4
t1.6 Creativity – 1	7	2

156 **Results**

157 **Group analyses**

158 **Accuracy**

159 To test effects of difficulty and task practice (defined across
 160 four chunks of 50-trial sub-blocks) on task accuracy, a
 161 repeated-measures ANOVA was carried out, with number of
 162 attributes (three/four/five) and sub-block as within-subject
 163 factors. As shown in Fig. 1 (solid lines), there was a main
 164 effect of difficulty ($F(2,50)=37.39, p<.001$). As the number
 165 of attributes increased accuracy dropped from 90% for the
 166 three attributes to 86% for four attributes and 84% for five
 167 attributes. No main effect of sub-block emerged
 168 ($F(3,75)=1.16, p=.329$), indicating that extensive learning is
 169 not necessary. For each number of attributes, the task-
 170 accuracy was higher than a bound obtained from an error-
 171 less version of the *TTB* heuristic (dashed lines),
 172 $F(1,25)=74.29, p<.001$.

173 **Reaction times**

174 Mean RT and the average number of trials (out of 200) in
 175 which the time limit was missed are given in Table 2.

176 A repeated-measures ANOVA, with number of attributes
 177 (three, four, five) as within-subject factor, revealed no effect of
 178 number of attributes, $F(1.25, 31.23)=0.056, p=.865$. Thus,
 179 although the task's difficulty increased and more information
 180 had to be considered, participants did not require more time.
 181 The number of trials in which the time limit was missed was
 182 only 0.6% of all trials and the average decision time was of
 183 around 1.5 s (see also Glöckner & Betsch, 2008, for similar
 184 results).

185 We also examined the correlation between RT and task
 186 accuracy. For each participant, we calculated the correlations
 187 between RT and accuracy, across all trials, separately for each
 188 number of attributes. The resulting mean correlations were
 189 negative, as predicted by the automatic *WADD* mechanism:
 190 $r = -0.199$ for three attributes, $r = -0.203$ for four attributes and
 191 $r = -0.172$ for five attributes. Interestingly, the negative

correlations between accuracy and RT remained even after
 controlling for the trials' difficulty: the partial correlation be-
 tween RT and accuracy remained negative: $r = -0.102$ for three
 attributes, $r = -0.084$ for four attributes, and $r = -0.062$ for five
 attributes. While these partial correlations are small, they are
 all significantly different from 0 (all p 's < .05, tested using a
 bootstrap procedure with 10,000 resamples) and in the same
 direction, suggesting that taking longer to decide reduces
 accuracy.

Strategy classifications

We next examined individual differences in decision strate-
 gies. We tested three potential strategies that participants could
 use when performing the task: the weighted average *WADD*,
TTB, and the Equal Weights rule combined with the *TTB*
 heuristic (*EQW-TTB*). According to the *EQW-TTB* strategy,
 one chooses the alternative for which the *non-weighted* aver-
 age is highest (i.e., the subjects average the values but ignore
 the importance weights). In cases where the non-weighted
 average for both alternatives is the same, participants choose
 according to the *TTB* heuristic, thus its name – *EQW-TTB*.¹
 We start using a simplified “trembling-hand” approach
 (Bröder, 2010), according to which the subject has a probabili-
 ty p to mistakenly report an alternative not predicted by the
 choice strategy. We use this approach in order to obtain an
 upper bound on the proportion of *TTB* use (we defer to the
 computational section, where we examine a more refined type
 of strategy-classifications allowing probabilistic errors and
 strategy mixtures).

The classifications were done separately for three/four/five
 attributes to test whether increased difficulty leads to more
 reliance on non-normative strategies. To classify the partici-
 pants based on the three strategies – *WADD*, *TTB*, and *EQW-*
TTB – we computed the probability of the data (200 choices)
 for each strategy and we selected the strategy that has the
 highest probability; see [Supplementary Material](#) for details
 of the classification procedure. The classification results are
 shown in Table 3 (see Table S1 in [Supplementary Material](#) for
 individual classifications and Tables S2, S3, and S4 for the
 normalized probabilities of the three strategies); 82% of the
 classifications (64 out of 78) are associated with normalized
 probability larger than .99, and 88% (69 out of 78) with a
 probability larger than .90.

As shown in Table 3, the majority of classifications (46 out
 of 78 in total, ~59%) belong to the compensatory (normative)
WADD strategy and another 8% (six in total) were a less op-
 timal but still compensatory *EQW-TTB* strategy. Only 29% of
 the classifications (23 in total) fell into the *non-compensatory*

¹ We also tested a less restrictive *Take-Two* heuristics (Dieckmann & Rieskamp, 2007), but none of the data sets were classified to this strategy and therefore we do not include it in Table 3.

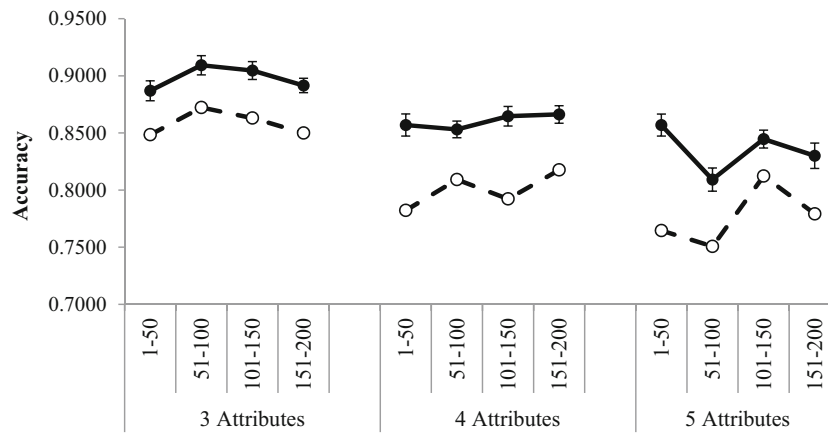


Fig. 1 Task-accuracy. Solid-lines: accuracy as a function of the number of attributes and of trial-number (in 50-trial blocks). Dashed lines: theoretical performance of *TTB* heuristic. Error bars represent within-subject standard errors (Cousineau, 2005)

239 *TTB* category. The amount of *WADD* classifications did not
 240 vary with the number of attributes. A summary of participants'
 241 accuracy as a function of strategy is shown in Fig. S1 (see
 242 Supplementary Material). We find that users of the *WADD*
 243 strategy had higher accuracy than *TTB* users, $t(67)=3.08$,
 244 $p=.003$. As reported in the Supplementary Material, this is
 245 not due to a speed-accuracy tradeoff.

246 **Attributes' weights**

247 Using logistic regression, we computed the subjective weights
 248 each participant gave to each of the attributes. Figure 2 shows
 249 these subjective weights for the three strategy subgroups (see
 250 Fig. S2 in the Supplementary Material for the group weights).
 251 These weights indicate that *WADD* users are better calibrated
 252 with the objective weights, the results of the *TTB* users show a
 253 strong overestimation of the most important attribute,
 254 confirming their reliance on a single dimension. A repeated-
 255 measures ANOVA on the attributes' weights of users of the
 256 *TTB* and *WADD* strategies revealed an interaction between the
 257 strategy used and the attributes' weights, for every number of
 258 attributes – for three attributes: $F(2,42)=21.36$, $p<.001$, for
 259 four attributes: $F(3,63)=13.93$, $p<.001$, and for five attributes:
 260 $F(4,92)=15.40$, $p<.001$. The *EQW-TTB* users showed the flat-
 261 test curves, consistent with the equality of weights character-
 262 izing this strategy.

263 **Reaction times: strategies and correlation with accuracy**

264 The *WADD* and the *TTB* strategies differ in their predic-
 265 tions concerning RT (Glöckner & Betsch, 2008). While
 266 according to *TTB* the RT should depend on whether there
 267 is a tie on the most important dimension, according to
 268 *WADD* the RT should depend on the absolute difference
 269 in the alternatives' weighted averages (ΔWA). To test this
 270 prediction, we applied to the log-RT-data of each partici-
 271 pant (we used log-RT in order to normalize the otherwise
 272 skewed values in the RT-distribution that may involve
 273 outliers; see also Glöckner & Betsch, 2008) a multiple
 274 linear regression with two factors: (i) ΔWA , (ii) a binary
 275 predictor of a tie on the most important dimension (i.e.,
 276 the most important attribute; see Table 1 for an example
 277 on which the tie variable equals 1 and $\Delta WA=0.9$). We
 278 compared the standardized regression coefficients for the
 279 participants who were classified as *WADD* users and those
 280 classified as *TTB* users. As predicted, the difficulty coef-
 281 ficient was stronger for the *WADD* users ($M = -0.43$, $SD =$
 282 0.12 ; $t(67)=3.56$, $p<.001$; Fig. 3, left), while the tie coef-
 283 ficient was higher in magnitude for the *TTB* users ($M =$
 284 0.06 , $SD = 0.14$) compared with the *WADD* users ($M = -$
 285 0.02 , $SD = 0.09$; $t(31.8)=2.31$, $p=.027$; Fig. 3, right).
 286 Unlike for the *TTB* users, for the *WADD* users the tie-
 287 coefficient was not significantly different from zero.
 288

289 **Computational-models of strategy choice:**
 290 **beyond the trembling hand**

291 While the trembling-hand classifications appear to have some
 292 validity, as they are supported by differences in subjective
 293 weights (Fig. 2) and in RTs (Fig. 3), there are a number of
 294 reasons to suspect that these classifications are a simplification
 295 and that the participants vary in a more continual, non-dichot-
 296 omous, manner. First, the weights are subject to individual

t2.1 **Table 2** Mean reaction time (RT) (standard deviations in parentheses)
 and average number of trials in which the time limit was missed (out of
 200), for each number of attributes

t2.2		Three attributes	Four attributes	Five attributes
t2.3	Mean-RT (SD)	1577 (512)	1593 (501)	1565 (519)
t2.4	Number of trials exceeding deadline	0.85	1.27	1.50

t3.1 **Table 3** Number of participants classified as users of each one of the three strategies (*WADD*, *TTB*, *EQW-TTB*) as a function of the number of attributes

t3.2 Strategy	t3.3 Number of attributes		
	Three	Four	Five
t3.4 WADD	15	14	17
t3.5 TTB	7	9	7
t3.6 EQW-TTB	3	2	1
t3.7 WADD/TTB*	1	1	1

*WADD/TTB** represents cases in which these two strategies had equal probabilities

297 variability even within a strategy group, indicating continuity
 298 rather than dichotomous strategies. Second, even for partici-
 299 pants classified as *TTB*, we obtain a *WADD* component in the
 300 RT-regression (Fig. 3). Finally, as recently discussed by Hilbig
 301 and Moshagen (2014), the trembling-hand type of error is not
 302 well matched with the natural assumptions of a *WADD* model,
 303 according to which choice-problems with lower ΔWA are ex-
 304 pected to have more errors. In order to extend the strategy
 305 classification to address these issues, we examined a number
 306 of computational models and carried out model-comparison
 307 using the aggregate Akaike Information Criterion (AIC;
 308 Akaike, 1973). As there are only six EQW classifications in
 309 our data (out of 78), we discard these and focus on contrasting
 310 between *WADD* and *TTB*.

311 Three new models were examined: (i) A probabilistic
 312 model, which in each trial deploys *WADD* with probability
 313 p and *TTB* with probability $(1-p)$. While this model as-
 314 sumes a trembling-hand error (as before), p provides a
 315 continuous measure of the degree of *WADD* use. (ii) A
 316 model that is like (i) with the exception that the *WADD*
 317 errors are not due to a trembling hand assumption, but
 318 rather are assumed to reflect Gaussian fluctuations in the
 319 *WADD* estimation (with zero mean and whose SD is a new
 320 model parameter); we still have a trembling hand param-
 321 eter for errors of the *TTB* heuristic (see Lee & Newell, 2011;

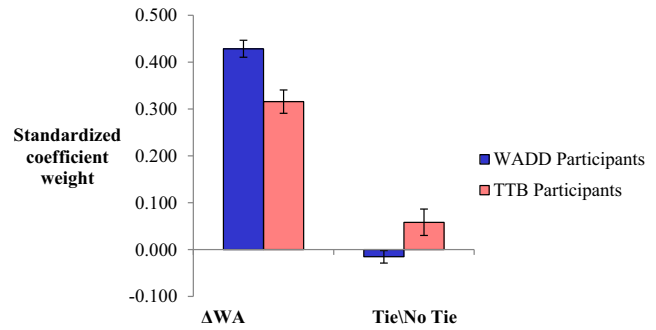


Fig. 3 Averaged standardized regression coefficients of: *left* (ΔWA – difficulty) – for graphic purposes we plotted the negatives of the difficulty-coefficients, and *right* (Tie/No Tie) – whether there was a tie on the most important dimension, separately for participants classified as users of *WADD* and *TTB* strategies. Error bars represent standard errors

Scheibehenne, Rieskamp, & Wagenmakers, 2013; for simi- 322
 323 lar approaches to mixture models in decision-making). 324
 325 (iii) A fully compensatory model, whose weights are char- 326
 327 acterized by a single parameter, α , based on normalized 328
 329 W_i^{α} , (where W_i are the normative weights (e.g., 4,3,2,1); 330
 331 note that $\alpha > 1$ results in an over-weighting of the high 332
 333 weights and under-weighting of the low weights, as in 334
 335 some version of the *PCS*; Glöckner et al., 2014). For each 336
 337 model, the parameters were fitted based on the probability 338
 339 of the data given the model (see Supplementary Material). 340
 341 The results are summarized at the group level in Table 4 342
 343 (see Supplementary Material for individual participants’ 344
 345 classifications).

335 We observe a clear picture. The single-strategy trembling- 336
 337 hand models provide the worst fit, followed by the compen- 338
 339 satory α -weight model, and then by the probabilistic strategy 340
 341 mixture *WADD/TTB* model. The best model by far is the mix- 342
 343 ture model with Gaussian *WADD* errors. Importantly, the pro- 344
 345 portion of *WADD* use in this model shows high consistency 346
 among the participants across the number of attributes (see 347
 348 Suppl.). Furthermore, we also find high correlations between 349
 the proportion of *WADD* use in the probabilistic model of 350
 individual subjects and subjective α -weights (all $|r|s > .75$; 351
 $p < .001$; see Supplementary Material). 352

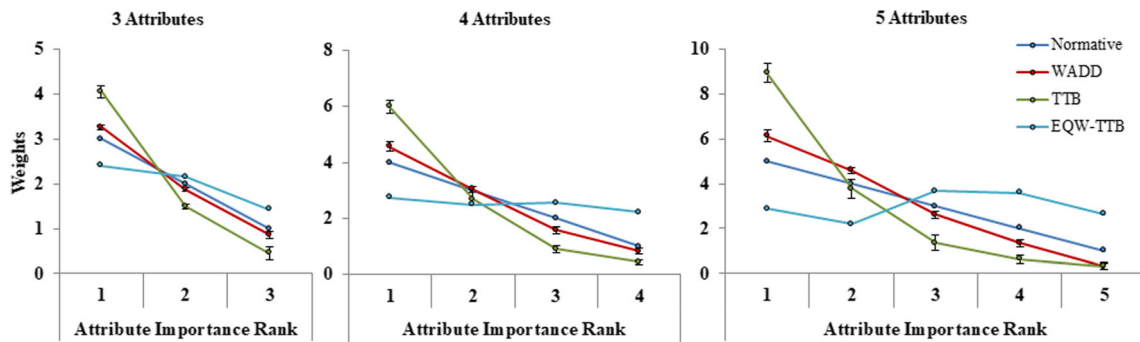


Fig. 2 Subjective weights for jobs with three (left), four (middle), and five (right) attributes, classified by strategy used. Error bars correspond to standard errors

t4.1 **Table 4** Akaike Information Criterion (AIC) table^a for models of strategy choice

t4.2 Model	Pure TTB	Pure WADD	Binary Mixture Model (TTB + WADD)	Gaussian mixture model (TTB + WADD)	α -weight model
t4.3 Three attributes	3,941	3,135	2,566	2,283	2,652
t4.4 Four attributes	4,413	3,914	3,028	2,681	3,241
t4.5 Five attributes	4,800	4,333	3,510	3,099	3,672

Note that AIC differences higher than 10 are considered decisive evidence (bold values indicate the best fits)

a The same conclusions are obtained with Bayesian Information Criterion (see Tables S8-S10 in Supplementary Material)

346 **Discussion**

347 In this study, we asked whether participants can rapidly
 348 carry out a complex weighted-averaging task. We used a
 349 job-interview framing and provided the participants with
 350 accuracy feedback. The number of attributes varied from
 351 three to five, a range exceeding the capacity of online
 352 analytical computations for speedy decisions. The results
 353 were surprising. First, the decision times (Mean-RT ~1.5
 354 s, which includes the visual encoding and the motor re-
 355 sponse) were much faster than the maximum allotted time,
 356 indicating reliance on intuitive gist-perceptions or heuris-
 357 tic rules (Saunders & Buehner, 2013). Second, despite the
 358 short RTs, the accuracy exceeded the bound that could be
 359 obtained on the basis of (error-less) non-compensatory
 360 strategies, such as *TTB*. Third, we found a negative cor-
 361 relation between accuracy and decision-time, consistent
 362 with evidence integration models, such as those based
 363 on the Decision-Field-Theory (Roe et al., 2001), the
 364 drift-diffusion model (Krajbich, Armel, & Rangel, 2010)
 365 or *PCS* (Glöckner et al., 2014). These results are consis-
 366 tent with those obtained by Glöckner and Betsch (2008,
 367 2012) in a multi-cue probabilistic inference task, and with
 368 their proposal of an automatic compensatory mechanism.

369 We examined two simplifying heuristics: One that (*TTB*) is
 370 non-compensatory, while the other (*EQW-TTB*) neglects the
 371 importance of the decision attributes. While the group-level
 372 performance exceeded the accuracy bound achievable from an
 373 error-free *TTB* heuristic, at the individual participants' level we
 374 found a certain amount of variability. The simplified
 375 (dichotomous) classification showed that, while most partici-
 376 pants relied on compensatory strategies, about 30% relied on
 377 the *TTB* heuristic. These participants were characterized by a
 378 peaked decision-weight pattern that overestimates the most
 379 important attribute (Fig. 2), by reduced task accuracy (without
 380 a Speed-Accuracy trade-off). *TTB*-users were also slower in
 381 trials with a tie on the most important dimension (Fig. 3, right
 382 panel). Unlike *TTB* users, most participants appeared to de-
 383 ploy a compensatory strategy that is likely to involve a noisy
 384 estimation of the weighted average (*WADD*; see also Glöckner
 385 & Betsch, 2008). The more refined (mixture) classifications
 386 indicate a continuum for participants' probability of deploying

a compensatory *WADD* strategy in each trial, ranging from a
 387 minimum of .23 to a maximum of 1. 388

389 We suggest that the presence of variability in decision strat-
 390 egies across the group reflects two potential ways of dealing
 391 with time pressure and information overload in decision mak-
 392 ing. The non-compensatory *TTB* heuristic is a lexicographic
 393 strategy that applies rules sequentially and neglects much of
 394 the information. Automatic and compensatory (*WADD*) strat-
 395 egies offer an alternative way to deal with information over-
 396 load. Instead of "calculating" the weighted average, these par-
 397 ticipants appear to carry out an "approximate" (noisy), but
 398 holistic estimation of it, consistent with an affective/intuitive
 399 decision mode (Kahneman, 2003). In particular, intuitive/
 400 holistic averaging is consistent with Kahneman's suggestion
 401 that intuitive processes are holistic in nature (see also
 402 Glöckner and Betsch, 2008) and are at the interface of percep-
 403 tion and cognition (Kahneman, 2003). This suggestion was
 404 also supported by recent empirical results showing dissocia-
 405 tions between intuitive and analytical averaging based on load
 406 manipulations (Rusou et al., 2017).

407 A potential mechanism to perform noisy weighted averag-
 408 ing estimations is Glöckner and colleagues' *PCS* model
 409 (Glöckner, Hilbig & Jekel, 2014). According to this model,
 410 the weighted average is computed in a neural network, which
 411 multiplies a values-vector with an importance-weights matrix.
 412 As our task involves some practice, the assumption that the
 413 decision mechanism includes learned weights (reflecting the
 414 attributes' importance) is not implausible.² Alternatively, the
 415 mechanism of weighted averaging could be mediated by a
 416 population code model (Brezis et al., 2016; Brezis et al.,
 417 2015), which operates using numerosity detectors (Dehaene,
 418 Molko, Cohen & Wilson, 2004; Piazza, Izard, Pinel, Le
 419 Bihan, & Dehaene, 2004). Future research is needed to probe

² This does not require to endorse all the assumptions of the *PCS* model, such as RT being based on convergence to asymptotic activation; an alternative assumption is an integration to boundary. The property of *PCS* that is important to our results is the parallel integration of values from all attributes. A somewhat similar approach is the accumulator model proposed by Lee and Cummins (2004), according to which the integrated values are subject to a response-boundary. This model, however, assumes that the values are integrated sequentially (in order of importance) and accounts for *TTB* use for low boundary values. Since in our data *TTB* users were not faster than *WADD* users, we support the parallel rather than the sequential integration of values.

420 the nature of the individual differences underlying the reliance
 421 on sequential and holistic processing in multi-attribute
 422 decisions.

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